#### STOCKHOLMS UNIVERSITET, MATEMATISKA INSTITUTIONEN, Avd. Matematisk statistik

# Suggested solutions: Econometric methods (MT5014) 2019-01-16 9-14

# Problem 1

(A) A time series  $\{r_t\}$  is weakly stationary if neither  $E(r_t)$  nor  $Cov(r_t, r_{t-l})$  depends on t (where l is an arbitrary integer); see also Tsay p. 30.

(B) Since the coefficient in front of  $r_{t-1}$  is 0.4, which satisfies |0.4| < 1, this time-series is weakly stationary (see Tsay p. 39). Hence by taking expectation, i.e  $E(r_t) = E(0.1 + 0.4r_{t-1} + a_t)$ , and using weak stationary and  $E(a_t) = 0$  we find  $E(r_t) = 0.1 + 0.4E(r_t)$ . Solving for  $E(r_t)$  gives  $E(r_t) = 0.1/0.6 = 1/6$ .

Similarly, we find  $V(r_t) = 0.4^2 V(r_{t-1}) + V(a_t)$ . Using weak stationary and  $V(a_t) = 1$  we find  $V(r_t) = 1/(1 - 0.4^2)$ .

(C) Consider the AR(1) time-series in (B) but change the coefficient 0.4 to 1 and the coefficient 0.1 to 0, then you have a definition of random walk; see also Tsay p. 72.

#### Problem 2

This relies on Tsay around p. 54-56.

(A) The forecast is  $\hat{r}_3(1) = 0.3 + 0.8r_3 = 0.3 + 0.8 \cdot 0.2 = 0.46$ .

(B) The forecast is  $\hat{r}_3(2) = 0.3 + 0.8\hat{r}_3(1) = 0.3 + 0.8 \cdot 0.46 = 0.668$ . The forecast error is,

$$e_3(2) = a_5 + 0.8a_4.$$

Hence, the associated variance is  $V(e_3(2)) = V(a_5) + V(0.8a_5) = (1+0.8^2)V(a_t) = 0.164.$ 

(C) The 1 step ahead forecast is,

$$\hat{r}_3(1) = 0.3 + 0.8r_3.$$

The 2 step ahead forecast is

$$\hat{r}_3(2) = 0.3 + 0.8\hat{r}_3(1) = 0.3 + 0.8(0.3 + 0.8r_3).$$

The 3 step ahead forecast is

$$\hat{r}_3(3) = 0.3 + 0.8\hat{r}_3(2) = 0.3 + 0.8(0.3 + 0.8(0.3 + 0.8r_3)).$$

The 4 step ahead forecast is

$$\hat{r}_4(3) = 0.3 + 0.8\hat{r}_3(3) = 0.3 + 0.8(0.3 + 0.8(0.3 + 0.8(0.3 + 0.8r_3))).$$

Continuing in this way clearly gives

$$\lim_{n \to \infty} \hat{r}_3(n) = 0.3(1 + 0.8 + 0.8^2 + \dots)$$

Recall that  $(1 + 0.8 + 0.8^2 + ...) = 1/(1 - 0.8)$ . Recall the general formula for  $E(r_t)$  for a weakly convering AR(1), which in this particular case gives  $E(r_t) = 0.3/(1 - 0.8)$ . It follows that  $\lim_{n\to\infty} \hat{r}_3(n) = E(r_t)$  and we are done.

#### Problem 3

We use the Chow test of structural change (p. 76-77 in the compendium). The F-statistic is

$$F = \frac{(18.3 - (9.1 + 4.3))/3}{(9.1 + 4.3)/(100 - 2 * 3)} = 11.4577.$$

Since the F-statistic is this large we can reject the hypothesis that the parameters are the same in the two sub-periods at a very small significance level (the exact significance level at which the hypothesis can be rejected can be looked up in a table for the F distribution). See p. 76-77 in the compendium for further details.

### Problem 4

(A)  $a_t$  should be a sequence of serially uncorrelated random vectors with mean zero and a covariance matrix  $\Sigma$  (which is typically positive definite) – Tsay p. 399.

(B) It is always weakly stationary (due to the assumption that the covariance matrix  $\Sigma$  exists) – Tsay p. 417-418.

- (C) See Tsay p. 419.
- (D) A feedback relationship Tsay p. 419.

### Problem 5

(A) Heteroskedasticity.

(B) It is unbiased. This can be seen by using the formula for the OLS estimate and the assumption that the expected values of the errors  $\tilde{\varepsilon}_j$  are zero. Hint: see

the last equation on p. 70 in the compendium.

(C) The GLS formula is (p. 89 in the compendium)

$$\hat{\boldsymbol{\beta}}_{GMM} = (\boldsymbol{X}^T \boldsymbol{\Omega}^{-1} \boldsymbol{X})^{-1} \boldsymbol{X}^T \boldsymbol{\Omega}^{-1} \boldsymbol{Y}$$

where for the model in the problem we have  $\boldsymbol{\Omega} = \text{diag}(e^{2X_1},...,e^{2X_n})$ , and

$$\boldsymbol{X} = \begin{bmatrix} 1 & X_1 \\ \vdots & \vdots \\ 1 & X_n \end{bmatrix}$$

$$\boldsymbol{Y} = \begin{bmatrix} Y_1 \\ \vdots \\ Y_n \end{bmatrix}$$

(D) It is BLUE (p. 90 in the compendium).

# Problem 6

(A) It is not unbiased. To see this use that

$$\hat{\boldsymbol{\beta}}_{OLS} = \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T \mathbf{Y} = \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T (\mathbf{X} \boldsymbol{\beta} + \boldsymbol{\varepsilon}) = \boldsymbol{\beta} + \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T \boldsymbol{\varepsilon}$$

and that we have not assumed that  $E[\boldsymbol{\varepsilon}|\mathbf{X}]$  is zero.

(B) See p. 116 in the compendium.