

Brownian motion

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Symmetric random walk

Consider a **symmetric random walk** $\{X_n, n \in \mathbb{N}\}$, which at each time is equally likely to go one step up or down, i.e., the X_n 's are i.i.d. with

$$\mathbb{P}(X_n = 1) = \mathbb{P}(X_n = -1) = \frac{1}{2}.$$

This is a Markov chain with $P_{i,i+1} = \frac{1}{2} = P_{i,i-1}$ for $i = 0, \pm 1, \pm 2, \dots$

- $S_n = \sum_{i=1}^n X_i$.
- $\mathbb{E}[S_n] = 0$ and $\text{Var}(S_n) = n$.
- By the CLT $\frac{1}{\sqrt{n}}S_n$ converges in distribution to a $\mathcal{N}(0, 1)$.

Scaling limit of a random walk

Let's **speed up this process** by taking smaller and smaller steps in smaller and smaller time intervals. If at each Δt time unit we take a step of size Δx up or down with equal probabilities, then the position at time t is

$$X(t) = \Delta x (X_1 + \cdots + X_{\lfloor t/\Delta t \rfloor}),$$

where the X_i 's are i.i.d. with $\mathbb{P}(X_i = 1) = \mathbb{P}(X_i = -1) = \frac{1}{2}$.

- $\mathbb{E}[X(t)] = 0$ and $\text{Var}(X(t)) = (\Delta x)^2 \lfloor \frac{t}{\Delta t} \rfloor$.

Go to the limit: let $\Delta x = \sigma\sqrt{\Delta t}$ for $\sigma > 0$, and let $\Delta t \rightarrow 0$.

- $\mathbb{E}[X(t)] = 0$ and $\text{Var}(X(t)) \rightarrow \sigma^2 t$.
- By the CLT $X(t)$ converges in distribution to a $\mathcal{N}(0, \sigma^2 t)$.

Since the random walk has independent and stationary increments, we expect the limiting process to have them as well.

Brownian motion

A stochastic process $\{X(t), t \geq 0\}$ is said to be a **Brownian motion** (BM) if

- (i) $X(0) = 0$;
 - (ii) $\{X(t), t \geq 0\}$ has independent and stationary increments;
 - (iii) for every $t > 0$, $X(t) \sim \mathcal{N}(0, \sigma^2 t)$.
-
- If $X(0) = x$, then $\{X(t) - x, t \geq 0\}$ is a BM.
 - If $\sigma = 1$, **standard Brownian motion** (SBM) $B(t) \sim \mathcal{N}(0, t)$. Any BM can be converted to the standard process by letting $B(t) = X(t)/\sigma$. From now on, we assume $\sigma = 1$.
 - By symmetry, $\{X(t), t \geq 0\}$ is distributed as $\{-X(t), t \geq 0\}$.

History

- 1827: the English botanist Robert **Brown** studied the motion of a small particle that is totally immersed in a liquid or gas.
- 1905: **Einstein** explained the process by assuming that the immersed particle was continually being subjected to bombardment by the molecules of the surrounding medium.
- 1918: **Wiener** gave the precise mathematical definition (also called Wiener process)

Continuous but not differentiable

Consider the SBM $\{X(t), t \geq 0\}$.

- $X(t)$ is a **continuous function** of t .

Intuition. We must show that $\lim_{h \rightarrow 0} (X(t+h) - X(t)) = 0$ a.s..

Note that the r.v. $X(t+h) - X(t) \sim \mathcal{N}(0, h)$ has mean 0 and variance h , and so it would seem to converge to a r.v. with mean 0 and variance 0 as $h \rightarrow 0$. □

- $X(t)$ is **nowhere differentiable**.

Intuition. Note that $\frac{X(t+h) - X(t)}{h} \sim h^{-1} \mathcal{N}(0, h) \sim \mathcal{N}(0, h^{-1})$ has mean 0 and variance $1/h$, which converges to ∞ if $h \rightarrow 0$. Hence, it is not differentiable. □

Brownian bridge

For $0 < s < t$, we are interested in the conditional distribution of $X(s)$ given that $X(t) = B$. The conditional density is

$$\begin{aligned}
 f_{X(s)|X(t)=B}(x) &= \frac{f_{X(s),X(t)}(x, B)}{f_{X(t)}(B)} = \frac{f_{X(s)}(x)f_{X(t-s)}(B-x)}{f_{X(t)}(B)} \\
 &= \frac{\frac{1}{\sqrt{2\pi s}} e^{-\frac{x^2}{2s}} \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{(B-x)^2}{2(t-s)}}}{\frac{1}{\sqrt{2\pi t}} e^{-\frac{B^2}{2t}}} = \frac{e^{-\left(\frac{x^2}{2s} + \frac{(B-x)^2}{2(t-s)} - \frac{B^2}{2t}\right)}}{\sqrt{2\pi \frac{s(t-s)}{t}}} \\
 &= \frac{e^{-\frac{t(t-s)x^2 + st(B^2 - 2Bx + x^2) - s(t-s)B^2}{2st(t-s)}}}{\sqrt{2\pi \frac{s(t-s)}{t}}} = \frac{e^{-\frac{(x-Bs/t)^2}{2s(t-s)/t}}}{\sqrt{2\pi \frac{s(t-s)}{t}}},
 \end{aligned}$$

hence it is normal with

$$\mathbb{E}[X(s) | X(t) = B] = \frac{s}{t}B \quad \text{and} \quad \text{Var}(X(s) | X(t) = B) = \frac{s}{t}(t-s).$$

Example

Example 10.1: Bicycle race.

Hitting times

Let $T_a = \inf\{t \geq 0 : X(t) \geq a\}$ be the **hitting time** of barrier a . For $a > 0$, since

$$\begin{aligned}\mathbb{P}(X(t) \geq a) &= \mathbb{P}(X(t) \geq a | T_a \leq t) \mathbb{P}(T_a \leq t) + \mathbb{P}(X(t) \geq a | T_a > t) \mathbb{P}(T_a > t) \\ &= \frac{1}{2} \mathbb{P}(T_a \leq t) + 0 \mathbb{P}(T_a > t) \\ &= \frac{1}{2} \mathbb{P}(T_a \leq t),\end{aligned}$$

we have that (using $y = x/\sqrt{t}$)

$$\mathbb{P}(T_a \leq t) = 2\mathbb{P}(X(t) \geq a) = \frac{2}{\sqrt{2\pi t}} \int_a^\infty e^{-x^2/2t} dx = \frac{2}{\sqrt{2\pi}} \int_{a/\sqrt{t}}^\infty e^{-y^2/2} dy.$$

By symmetry, for $a < 0$ the distribution of T_a is the same as that of T_{-a} , hence we obtain $\mathbb{P}(T_a \leq t) = \frac{2}{\sqrt{2\pi}} \int_{|a|/\sqrt{t}}^\infty e^{-y^2/2} dy$. By computing the density $f_{T_a}(t)$, we get $\mathbb{E}(T_a) = \infty$.

The maximum of Brownian motion

Let $M(t) = \max_{0 \leq s \leq t} X(s)$ be the **maximum of Brownian motion**. For $a > 0$,

$$\mathbb{P}(M(t) \geq a) = \mathbb{P}(T_a \leq t) = 2\mathbb{P}(X(t) \geq a) = \frac{2}{\sqrt{2\pi}} \int_{a/\sqrt{t}}^{\infty} e^{-y^2/2} dy.$$

Moreover, by symmetry,

$$2\mathbb{P}(X(t) \geq a) = \mathbb{P}(X(t) \geq a) + \mathbb{P}(X(t) \leq -a) = \mathbb{P}(|X(t)| \geq a).$$

Hence $M(t)$ has the same distribution of $|X(t)|$.

The reflection principle

Theorem (The reflection principle)

Let $\{X(t), t \geq 0\}$ is a SBM and T a stopping time. The process $\{X_T(t), t \geq 0\}$ defined as

$$X_T(t) = \begin{cases} X(t), & 0 \leq t \leq T, \\ 2X(T) - X(t), & t > T, \end{cases}$$

is also a SBM.

Example: exam 2021.

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Variations on Brownian motion

- If $\{B(t), t \geq 0\}$ is a SBM and $\mu \in \mathbb{R}$, then the process $\{X(t) = B(t) + \mu t, t \geq 0\}$ is a **Brownian motion with drift** μ . Since $\frac{B(t)}{t} \sim \mathcal{N}(0, 1/t)$, and its variance converges to 0, we have that $\frac{B(t)}{t} \rightarrow 0$ in probability and a.s.. Hence, as $t \rightarrow \infty$,

$$\frac{X(t)}{t} = \frac{B(t)}{t} + \mu \rightarrow \mu \quad \text{a.s..}$$

- If $\{Y(t), t \geq 0\}$ is a BM with drift μ and variance parameter σ^2 , the the process $\{X(t), t \geq 0\}$, defined by $X(t) = e^{Y(t)}$ is a **geometric Brownian motion**.

Geometric Brownian motion

Given the history of the process up to time $s < t$, the expected value of the process at time t is

$$\begin{aligned}\mathbb{E}[X(t) | X(u), 0 \leq u \leq s] &= \mathbb{E}[e^{Y(t)} | Y(u), 0 \leq u \leq s] \\ &= \mathbb{E}[e^{Y(t)-Y(s)+Y(s)} | Y(u), 0 \leq u \leq s] \\ &= e^{Y(s)} \mathbb{E}[e^{Y(t)-Y(s)} | Y(u), 0 \leq u \leq s] \\ &= X(s) \mathbb{E}[e^{Y(t)-Y(s)}]\end{aligned}$$

and, since $Y(t) - Y(s) \sim \mathcal{N}(\mu(t-s), \sigma^2(t-s))$,

$$\mathbb{E}[e^{Y(t)-Y(s)}] = e^{\mathbb{E}[Y(t)-Y(s)] + \text{Var}(Y(t)-Y(s))/2} = e^{\mu(t-s) + \sigma^2(t-s)/2},$$

hence,

$$\mathbb{E}[X(t) | X(u), 0 \leq u \leq s] = X(s)e^{(t-s)(\mu + \sigma^2/2)}.$$

Similarly, we can compute

$$\text{Var}(X(t) | X(u), 0 \leq u \leq s) = (X(s))^2 e^{2\mu(t-s)} (e^{2\sigma^2(t-s)} - e^{\sigma^2(t-s)}).$$

Application

Geometric BM is useful in the **modeling of stock prices** over time when the percentage changes are i.i.d..

- Suppose that $X(n)$ is the price of some stock at time n and $Y(n) = X(n)/X(n-1)$, $n \geq 1$ are i.i.d..
- We have that $X(n) = Y(n)X(n-1) = Y(n)Y(n-1)\dots Y(1)X(0)$.
- Then $\log(X(n)) = \sum_{i=1}^n \log(Y(i)) + \log(X(0))$ and, since the $\log(Y(i))$'s are i.i.d., $\{\log(X(n)), n \geq 0\}$ is approximately (for large n) a BM with a drift, so $\{X(n), n \geq 0\}$ is approximately a geometric BM.

The maximum of Brownian motion with drift

Let $\{X(t), t \geq 0\}$ be a BM with drift μ and variance parameter σ^2 , and define

$$M(t) = \max_{0 \leq s \leq t} X(s)$$

to be the **maximal value of the process** up to time t . We are interested in the distribution of $M(t)$.

- For $y > x$, we have that $\mathbb{P}(M(t) \geq y | X(t) = x) = e^{-\frac{2y(y-x)}{t\sigma^2}}$, $y \geq 0$.
The proof uses the fact that the conditional distribution of $X(s), 0 \leq s \leq t$ given $X(t)$ does not depend on μ .
- Then we get $\mathbb{P}(M(t) \geq y) = e^{2y\mu/\sigma^2} \bar{\phi}\left(\frac{y+\mu t}{\sigma\sqrt{t}}\right) + \bar{\phi}\left(\frac{y-\mu t}{\sigma\sqrt{t}}\right)$, with $\bar{\phi}(x) = 1 - \phi(x) = \mathbb{P}(Z > x)$ for $Z \sim \mathcal{N}(0, 1)$.
- For $y > 0$, recall that $M(t) \geq y$ iff $T_y \leq t$.

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Gaussian processes

A stochastic process $\{X(t), t \geq 0\}$ is called a **Gaussian process** (or normal process) if $X(t_1), \dots, X(t_n)$ have a multivariate normal distribution for all $t_1, \dots, t_n, n \geq 1$.

- Let Z_1, \dots, Z_m be i.i.d. $\mathcal{N}(0, 1)$ and let

$$X_i = \sum a_{ij} Z_j + \mu_i$$

for some constants $a_{ij}, 1 \leq i \leq n, 1 \leq j \leq m$ and $\mu_i, 1 \leq i \leq n$.

Then the r.v.'s X_1, \dots, X_n have a multivariate normal distribution.

- A multivariate normal distribution is completely determined by the marginal mean values and the covariance matrix.

Brownian motion is Gaussian

- If $\{X(t), t \geq 0\}$ is a BM, since each $X(t_1), \dots, X(t_n)$ can be expressed as a linear combination of the independent normal variables $X(t_1), X(t_2) - X(t_1), \dots, X(t_n) - X(t_{n-1})$ it follows that **BM is a Gaussian process**.
- Hence SBM can also be defined as a Gaussian process with $\mathbb{E}[X(t)] = 0$ and $\text{Cov}(X(s), X(t)) = \min\{s, t\}$. Indeed, for $s \leq t$

$$\begin{aligned}\text{Cov}(X(s), X(t)) &= \text{Cov}(X(s), X(s) + X(t) - X(s)) \\ &= \text{Cov}(X(s), X(s)) + \text{Cov}(X(s), X(t) - X(s)) \\ &= \text{Cov}(X(s), X(s)) = \text{Var}(X(s)) = s.\end{aligned}$$

The Brownian bridge is Gaussian

- Let $\{X(t), 0 \leq t \leq 1 \mid X(1) = 0\}$ be a standard Brownian bridge (SBB). Since in general $X(t) \mid X(1) = B \sim \mathcal{N}(tB, t(1-t))$, we have that $X(t) \mid X(1) = 0 \sim \mathcal{N}(0, t(1-t))$.
- The conditional distribution is a multivariate normal, hence the standard **Brownian bridge is a Gaussian process** with mean value 0 and covariance $\text{Cov}(X(s), X(t)) = s(1-t)$, $s \leq t$. Indeed, since $\mathbb{E}[X(u) \mid X(1) = 0] = 0$ for all $u < 1$,

$$\begin{aligned}\text{Cov}(X(s), X(t) \mid X(1) = 0) &= \mathbb{E}[X(s)X(t) \mid X(1) = 0] \\ &= \mathbb{E}[\mathbb{E}[X(s)X(t) \mid X(t), X(1) = 0] \mid X(1) = 0] \\ &= \mathbb{E}[X(t)\mathbb{E}[X(s) \mid X(t)] \mid X(1) = 0] = \mathbb{E}[X(t)\frac{s}{t}X(t) \mid X(1) = 0] \\ &= \frac{s}{t}\mathbb{E}[(X(t))^2 \mid X(1) = 0] = \frac{s}{t}t(1-t) = s(1-t).\end{aligned}$$

- Alternative definition: if $\{X(t), t \geq 0\}$ is a SBM, then $\{Z(t) = X(t) - tX(1), 0 \leq t \leq 1\}$ is a SBB.

Integrated Brownian motion

If $\{X(t), t \geq 0\}$ is a BM, the process $\{Z(t) = \int_0^t X(s) ds, t \geq 0\}$ is called **integrated Brownian motion**.

- It can be used to model the price of a commodity throughout time. Let $Z(t)$ be the price at time t and assume the rate of change $X(t) = \frac{d}{dt}Z(t)$ follows a BM. Then $Z(t) = Z(0) + \int_0^t X(s) ds$.
- Since BM is a Gaussian process, also **the integrated BM is a Gaussian process**. When $\{X(t), t \geq 0\}$ is a SBM, we get $\mathbb{E}[Z(t)] = \mathbb{E}[\int_0^t X(s) ds] = \int_0^t \mathbb{E}[X(s)] ds = 0$ and, for $s \leq t$,

$$\begin{aligned}\text{Cov}(Z(s), Z(t)) &= \mathbb{E}[Z(s)Z(t)] = \mathbb{E}\left[\int_0^s \int_0^t X(u)X(v) dudv\right] \\ &= \int_0^s \int_0^t \mathbb{E}[X(u)X(v)] dudv = \int_0^s \int_0^t \min\{u, v\} dudv \\ &= \int_0^s \left(\int_0^u v dv + \int_u^t u dv\right) du = s^2 \left(\frac{t}{2} - \frac{s}{6}\right).\end{aligned}$$

Stationary processes

A stochastic process $\{X(t), t \geq 0\}$ is a **stationary process** if for all $n \geq 1$ and $s > 0$ the random vectors $(X(t_1), \dots, X(t_n))$ and $(X(s + t_1), \dots, X(s + t_n))$ have the same joint distribution.

Examples:

- An ergodic continuous-time Markov chain $\{X(t), t \geq 0\}$ when $\mathbb{P}(X(0) = j) = P_j$ for each state j , i.e., when the initial state is chosen according to the limiting probabilities.
- The process $\{X(t) = N(t + L) - N(t), t \geq 0\}$ with $L > 0$ constant and $\{N(t), t \geq 0\}$ Poisson Process with rate λ . It follows from the stationary and independent increment assumption of Poisson processes.
- BM is not a stationary process.

Example 10.5: the random telegraph signal process.

Weakly stationary processes

A stochastic process $\{X(t), t \geq 0\}$ is a **weakly stationary process** if for all $s, t > 0$, $\mathbb{E}[X(t)] = c$ constant and $\text{Cov}(X(t), X(t+s)) = R(s)$ does not depend on t .

For Gaussian processes, weakly stationarity implies stationarity, because multivariate normal distributions are determined by their mean values and the covariance matrix. Note that BM is not weakly stationary.

Example 10.6: the Ornstein-Uhlenbeck process.

Example 10.8: the random telegraph signal process.

Exercises

Session 11. Chapter 10: 1-4, 6.

Session 12. Chapter 10: 7, 9, 10, 32, 35. Suggested extra: 11, 33, 34.

Note that answers might contain some integrals that are very hard to evaluate. It is okay to provide answers as not evaluated integrals.